

1 Does exposure to different types of menstrual protections
2 affect the vaginal environment?

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35 **Abstract**

36 The vaginal ecosystem is a key component of women’s health. It also represents an
37 ideal system for ecologists to investigate the consequence of perturbations on species
38 diversity and emerging properties between organisational levels. Here, we study how
39 exposure to different types of menstrual protection products is linked to microbial,
40 immunological, demographic, and behavioural measurements in a cohort of young
41 adult women who reported using more often tampons ($n = 107$) or menstrual cups
42 ($n = 31$). We first found that cup users were older and smoked less than tampon
43 users. When analysing health indicators, we detected potential associations between
44 cups use reporting fungal genital infection. Multivariate analysis confirmed that, in
45 our cohort, reporting using cups over tampons was associated with the higher odds
46 ratio to report a fungal genital infection diagnosis by a medical doctor within the last
47 3 months. We did not detect significant differences between groups in terms of their
48 bacterial vaginal microbiota composition and found marginal differences in the level
49 of expression of 20 cytokines. However, a multivariate analysis of these biological
50 data identified some level of clustering based on the type of menstrual production
51 preferred (cups or tampons). These results suggest that exposure to different types
52 of menstrual products could influence menstrual health. Larger studies and studies
53 with a more powered setting are needed to assess the robustness of these associations
54 and identify causal mechanisms.

55 **Keywords:** vaginal microbiota; fungal infections; immunity; epidemiology; women’s
56 health

57 Introduction

58 There is a long history of considering organisms as ecosystems, and *vice versa* [1]. This
59 parallel has been particularly exploited to study within-host parasite dynamics [2]. With
60 the advent of metagenomics, there is a renewed interest in applying ecological theory to
61 understand microbiota dynamics [3, 4]. Beyond parasites or even the microbiota, the
62 importance of encompassing the diversity of these ecosystems, e.g. including immune
63 cells, is increasingly acknowledged, especially since the border between self and non-self
64 is tenuous in this context [5].

65 Some general questions pertaining to ecosystem functioning can be studied by con-
66 sidering an individual host as an ecosystem. For instance, it is challenging to clearly
67 define properties that emerge at a higher level of organisation in large-scale ecologi-
68 cal systems. However, the definition and measurement of individual health outcomes
69 (i.e., systems-level outcomes) is routine in clinical research. In contrast to free-living
70 ecosystems, individual hosts are also smaller, more numerous, and readily replicable,
71 making empirical ecosystem ecology research more accessible and reproducible [6].

72 The vaginal ecosystem is an ideal study system to understand how within-host inter-
73 actions shape host health. First, the vaginal microbiota is a key component of women’s
74 health throughout their lives [7], modulating the risk of diseases such as fungal genital
75 infections, urinary tract infections, and sexually transmitted infections [8], such as human
76 immunodeficiency virus [9]. Contrary to its gut counterpart, the vaginal microbiota is
77 characterised by relatively few, clearly-defined, community state types (CST), most of
78 which are dominated by one species of *Lactobacillus* bacteria [10]. Second, the microbes
79 in the vaginal environment, including bacteria, fungi, and viruses, are involved in a tight
80 cross-talk with the local innate and adaptive immune response, the components of which
81 are known to be adapted to this highly variable environment [11]. In particular, the
82 interaction between the immune response and sex hormones throughout the menstrual
83 cycle is well documented [12]. Finally, the vaginal ecosystem is perturbed by factors
84 including menses, sexual activities, and antibiotic treatments [13]. Longitudinal follow-
85 ups demonstrate that such perturbation can lead to shifts from one CST to another [14].
86 Perturbations that jeopardise the balance between immunity, hormones, and the vaginal

87 microbiota could also hinder the efficiency of the immune system to control bacteria,
88 fungal, and viral pathogens [13].

89 Here, we explore whether exposure to a particular type of menstrual protection
90 product, namely menstrual cups or tampons, is associated with differences in the vaginal
91 environment. To this end, we analyse the vaginal microbiota, immune response (i.e.,
92 cytokines and antibodies), and, more generally, the emerging property of this ecosystem,
93 which revolves around menstrual health [15].

94 Among the variety of products used during menses, menstrual cups are perceived
95 as a safe, practical, economical, and ecologically friendly alternative to tampons and
96 sanitary pads [16]. The majority of women using them report wanting to continue to use
97 them, both in high and low-income settings, showing a good level of acceptability [17,
98 16]. Nevertheless, there have been some reported cases of toxic shock syndrome, renal
99 colic, and allergies linked with menstrual cup usage [16]. For example, higher levels of
100 *Staphylococcus aureus* growth have been reported in menstrual cups compared to tam-
101 pons [18]. In 2019, a systematic review identified 12 clinical trials and 23 observational
102 studies comparing this type of menstrual protection to others [16]. In the vast majority
103 of existing studies, the outcome studied was the practicality of the cups implementation,
104 e.g. acceptability or risk of leakage. Therefore, we know little about the potential mi-
105 crobiological and immunological implications of cup usage compared to other menstrual
106 hygiene products.

107 In the present study, we analysed biological, demographic, and behavioural longitudi-
108 nal data from 138 women. We compared two populations of women defined by the type
109 of menstrual protection they reported using most often (tampons or menstrual cups) and
110 analysed microbiological, immunological, and clinical data while statistically accounting
111 for demographic and behavioural differences. Using statistical modelling, we identified
112 profile differences depending on the type of menstrual product used.

113 **Materials and methods**

114 **Cohort description and data curation**

115 Women included in the present study were enrolled in the PAPCLEAR longitudinal
116 clinical study [19], which followed 149 women between 2016 and 2020 to study human

117 papillomavirus (HPV) infections [19]. The women enrolled were between 18 and 25 years
118 old, lived in the area of Montpellier (France), and reported at least one new partner within
119 the last 12 months. Their status for HPV and other genital infections, immunological
120 responses (antibodies and cytokines), and behaviours were followed for up to two years
121 (see [20] for details about the study protocol).

122 We selected participants who reported using tampons or cups as menstrual products
123 and for whom detailed cytokine profiles, microbiota metabarcoding data, and antibody
124 data at the inclusion visit were available. We assigned tampon or menstrual cup categories
125 when a participant reported using either type of menstrual product over 75% of the time
126 over the whole duration of the study (see Figure S1 for details on how the assignment was
127 performed). Women using menstrual cups more frequently than tampons are denoted
128 ‘menstrual cups users’ for convenience throughout the manuscript. There was no difference
129 in follow-up duration between women using mostly cups or mostly tampons (Table 1).
130 Other analyses of demographic, behavioural, and biological data were performed on the
131 first visit of the participant (inclusion visit). Overall, our analysis includes data from
132 138 women.

133 Since the PAPCLEAR study was designed to understand the ecology of HPVs and
134 their interaction with immune responses and vaginal microbiota, the data from this cross-
135 sectional study is not perfectly balanced in terms of exposure to each type of menstrual
136 product. However, given the current lack of studies, it greatly improves our understanding
137 of women’s menstrual health.

138 **Biological analyses**

139 HPV infections were assessed from cervical smears using the LiPA₂₅ genotyping assay,
140 which discriminates between 25 genotypes [21]. Further details on HPV detection and
141 prevalence in this cohort can be found elsewhere [20]. A ‘focal’ infection was defined as
142 an infection by the same HPV genotype at the first and second visits.

143 The microbiota metabarcoding was performed on 200 µL of vaginal swabs specimen
144 stored at -80° in Amies medium. DNA extraction was performed using the MagAt-
145 tract PowerMicrobiome DNA/RNA Kit (Qiagen). Next-generation sequencing of the
146 V3-V4 region of the 16S gene [22] was performed on an Illumina HiSeq 4000 platform

147 (150 bp paired-end mode) at the Genomic Resource Center at the University of Mary-
148 land School of Medicine. Taxonomic assignment was performed using the internal soft-
149 ware package SpeciateIT (<https://github.com/Ravel-Laboratory/speciateIT>) and
150 the community state type (CST) was determined using the VALENCIA software package
151 [23].

152 Antibodies were analyzed using a multiplex Luminex assay [24], as already described
153 in the context of HPV infections [20]. Cytokine data were obtained using MesoScale
154 discovery technology from vaginal secretions collected using ophthalmic sponges placed
155 directly on the cervical os for approximately one minute, as described in a previous
156 study which analyses this data in the context of HPV infections [25]. We used the same
157 protocol to obtain values that were normalised per total protein concentration in the
158 sample.

159 **Statistical analyses**

160 To study the difference between the main characteristics of our two populations of interest
161 (Table 1 and Table 2), we used χ^2 -test or Kruskal-Wallis rank sum tests when applicable.

162 We then performed multivariate analyses using generalised linear models (GLMs)
163 with a binomial distribution for the response variable for the models shown in Figure 1A,
164 Table 3, and Supplementary Tables S4 to S8. In the first series of models, the response
165 variable was the type of menstrual protection (cups or tampons). In the second series,
166 the response variable was being diagnosed or not with a fungal genital infection within
167 the last three months by a medical doctor.

168 We identified 15 covariates of interest (Table S1). Given the exploratory nature of
169 the study, we built models with all the possible combinations of covariates as explanatory
170 variables. We then selected the best models using the Akaike Information Criterion
171 corrected for small sample sizes (AICc). We used a lowest AICc +5 interval to identify
172 the most probable best models [26]. For each model, we computed the odds ratios
173 associated with each predictor along with a 95% confidence interval.

174 For clustering, factor analysis of mixed data was used when combining both factor and
175 numeric data (Figure 3B and S2), and multiple correspondence analysis was used when
176 analysing categorical data (Figures S2 and S4) [27]. Non-overlapping 95% confidence

177 ellipses indicate a statistically significant difference between cups and tampons users. The
178 covariates used for the socio-economic analysis using a multiple correspondence analysis
179 are listed in Table S2.

180 Statistical analyses were performed in R 4.1.3 [28] with the with the `kruskal.test`
181 function for Kruskal-Wallis rank sum tests (in Tables 1 and 2 for numeric data), `chisq.test`
182 for χ^2 test (in Table 1 and Figure 2B for proportional data), `wilcox.test` function for
183 Mann-Whitney test (Figure 2C) and the `glmulti` function for binomial regressions [29]
184 (in Figure 1A and Table 3). The factor analysis of mixed data and the multiple correspon-
185 dence analysis (in Figures 1B, 3B, S2, S3, and S4) were processed and represented using
186 the `FactorMineR` and the `FactoExtra` (<https://github.com/kassambara/factoextra>)
187 packages [27].

188 Results

189 The 138 women included in the analysis were aged from 18 to 25 years old and primarily
190 university students (119/138, 86%). We stratified the population according to the most
191 frequent type of menstrual product used, *i.e.* either tampons, $n = 107$, or menstrual
192 cups, $n = 31$. Analysis of the main demographic characteristics (shown in Table 1) using
193 a generalised linear model (GLM) selection approach identified significant differences
194 between these two populations in terms of age and smoking status: cups users were older
195 and reported smoking less than tampon users (Figure 1A and Table S3). On the other
196 hand, a multiple correspondence analysis on 12 covariates associated with socio-economic
197 status and listed in Table S2 did not show any difference between the two populations
198 (Figure 1B).

199 We then explored associations between the type of menstrual product used and
200 microbiological covariates. First, we analysed vaginal microbiota diversity using 16S
201 metabarcoding data (Figure 2A). We found no significant difference in community state
202 types (CST) composition (Figure 2B), although there were some qualitative differences.
203 For instance, among our menstrual cup users, we found no occurrence of CST II, which
204 is dominated by *L. gasseri*, but slightly more CST I and V, which are dominated by
205 *L. crispatus* and *L. jensenii* [10]. We also did not find any significant differences in
206 microbiota diversity, assessed using Shannon diversity index, between women using
207 menstrual cups or tampons (Figure 2C).

208 We then analysed six health-related covariates, most of which correspond to infections
209 (Table 2). The only significant difference was that women using cups also reported more
210 often having been diagnosed with fungal genital infections within the last 3 months. To
211 further investigate this association, we used our GLM selection method to identify the
212 covariates associated with this fungal genital infection. We found that reporting using
213 cups is the only predictor with a significant odds ratio (OR of 4.73, 95% CI: 1.44-16.1)
214 associated with the risk of reporting a fungal genital infection (the number of lifetime
215 partners is also present in the model, but not significant). Similar models for the other
216 five health-related covariates are shown in Appendix (Tables S4 to S8).

217 To investigate the potential effect of menstrual cups on the local immune response,

218 we analysed cytokine and chemokine relative concentrations in cervical samples. Among
219 the 20 analytes measured, IL-10, MIP-1 α and TNF- α appeared to be significantly lower
220 in women using cups (IL-10: $p = 0.012$, MIP-1 α : $p = 0.049$ and TNF- α : $p = 0.036$),
221 although these associations did not withstand correction for multiple testing comparisons
222 ($p = 0.24$, $p = 0.32$, and $p = 0.32$ respectively) (Figure 3A).

223 Finally, we performed a profile analysis using a factor analysis of mixed data approach
224 involving CST data, Shannon diversity index, and cytokine and chemokine relative con-
225 centrations. Women who use tampons and women using menstrual cups fall into two
226 largely non-overlapping clusters (Figure 3B and Supplementary Figure S2), suggesting
227 that the type of menstrual product used could be associated with a different immunolog-
228 ical and microbial environment. Conversely, a similar multiple correspondence analysis
229 approach using blood seropositivity status for immunoglobulins G (IgG) and immunoglob-
230 ulins M (IgMs) of several sexually transmitted infections, including HPVs, detected no
231 clustering (Supplementary Figure S3), hinting that the women in these two groups have
232 similar exposure risks to sexually transmitted infections.

233 Discussion

234 The vaginal ecosystem is an essential part of women’s health and several ‘perturbations’
235 such as menses, sexual intercourse, or drug treatment can cause pronounced shifts [7]. In
236 this study, we analysed whether exposure to different types of menstrual products, namely
237 menstrual cups or tampons, could be associated with differences in health indicators or,
238 more generally, in the vaginal environment.

239 Menstrual cups are gaining in popularity as an environmentally sustainable and
240 affordable type of menstrual protection, but the data surrounding their use are limited,
241 especially outside low-income countries [16]. Therefore, we first analysed the profiles of
242 the two populations of women defined based on the type of menstrual protection they
243 reported using most. In terms of demographic covariates, we found that the main factors
244 that were associated with preferential use of cups over tampons were age (cup users
245 were older) and smoking status (tampon users smoked more). In terms of socio-economic
246 status, a multiple correspondence analysis did not suggest any differences between the two
247 populations, which could be due to the fact that the majority of the cohort member was

248 university students. A similar multiple correspondence analysis based on serological data
249 for a variety of sexually transmitted infections also did not find any difference, further
250 suggesting that these two cohorts are comparable in terms of their general lifestyle.

251 On the microbiological side, we did not find a significant difference between the
252 bacterial community compositions of the vaginal environment depending on the type of
253 menstrual product used. A recent preprint [30] studied vaginal microbiota in a large
254 cohort of 3,345 women aged 18 to 98 with a median age of 32 years old. Their results
255 show that women who reported using menstrual cups are more often associated with a
256 *Lactobacillus crispatus* cluster (which also includes *L. jensenii*, and could be described
257 as encompassing both CST-I and CST-V); which is consistent with the trend detected
258 in our dataset.

259 Moving to health-related factors, we detected a strong association between reporting
260 using menstrual cups and a recent diagnosis of a fungal genital infection. None of the
261 other five health-related variables that we analyzed, such as reporting urinary tract
262 infections or the diagnosis of chlamydia infections, were significantly different. To further
263 investigate this trend, we used a generalised linear model selection approach. In almost
264 all the best models, reporting using cups more than tampons was significantly associated
265 with a higher risk of reporting fungal genital infections. It is noteworthy that the number
266 of partners was present in most of the best models (98.75% of the time), although never
267 significant.

268 Finally, we investigated the host component of the vaginal environment by analyzing
269 local cytokine and chemokine concentrations. Three cytokines or chemokines were found
270 to exhibit potentially lower concentrations in women using cups (IL-10, MIP-1 α , and
271 TNF- α), although significance did not withstand correction by multiple hypothesis testing.
272 Furthermore, the joint analysis of microbiota and immunological data suggests that
273 women tend to segregate into two clusters based on the type of menstrual product they
274 use most suggesting general profile differences in terms of their vaginal environment.

275 Our study has several limitations, the strongest being the relatively small sample size
276 of the cohort used ($n = 138$), which was originally designed to study HPV infections.
277 This may hinder our ability to detect moderate or subtle changes induced by menstrual
278 cups in the vaginal environment. For example, we included the use of lubricants in the

279 analysis, but more detailed studies could also include details on contraception methods or
280 probiotic use. Another limitation of the study lies in its cross-sectional nature. Further
281 longitudinal analyses would be helpful to establish long-term potential impacts of the
282 use of menstrual cups on the local environment, for example, on the vaginal microbiota
283 composition, which can be highly variable over time for some women [31]. In particular,
284 prospective or retrospective cohort studies, where participants are stratified based on
285 the type of menstrual products used, could be a powerful means to study the occurrence
286 of a particular negative outcome, i.e. here fungal infections.

287 Regarding these infections, we cannot provide any further details regarding the fungal
288 species or their abundance with the current data. A quantitative polymerase chain
289 reaction approach targeting the ITS locus coupled with a metagenomics approach could
290 provide a more accurate picture, especially thanks to existing gene catalogs for the vaginal
291 microbiota, such as VIRGO which already includes more than 10,000 fungal genes [32].

292 The existing literature on the usage of menstrual cups focuses on availability and
293 acceptability, yet its health implications remain understudied. A recent systematic
294 review reported nine cases of urinary tract complaints [16], which we did not detect as
295 being significantly more frequent among menstrual cups users than for tampons users.
296 Conversely, the primary adverse health outcome associated with menstrual cup use in our
297 study, i.e. fungal infections, was not included in the systematic review. Our results are
298 consistent with another recent study, which does not show adverse effects of menstrual
299 cups on the vaginal microbiota, and even potentially an increased proportion of some
300 *Lactobacillus*-dominated communities, which are considered beneficial for maintaining
301 the vaginal health [30]. These findings can help inform public health policies regarding
302 the use of menstrual cups and underline the need for larger, balanced cohort studies.

303 Health is partly the result of dynamic interactions between hosts and their microbes
304 that have a long coevolutionary history. The human vaginal microbiota is an ideal
305 system for studying these questions for several reasons. First, its composition is a known
306 health moderator — *Lactobacillus*-dominated communities tend to decrease the risk of
307 acquisition of sexually transmitted infections [10]. Second, the composition of the vaginal
308 microbiota can vary over time as a result of menstrual cycles or age, as well as a variety
309 of ‘perturbations’ including sexual intercourse and chemical interventions [14]. These

310 two points are not unique to the vaginal microbiota but contrarily to other microbiota,
311 especially in the gut, the diversity is limited (both between and within women) and
312 the variability better understood. From an evolutionary point of view, human vaginal
313 microbiota are unique for the high prevalence of *Lactobacillus*-dominance: Lactobacilli
314 are rarely dominant in other mammals, including non-human primates [33]. This human
315 specificity complicates experimental studies using animal model systems. However, three-
316 dimensional cell culture models now offer new opportunities for manipulative experiments,
317 for instance to test how the microbiota composition affects the interaction between a
318 parasite and host cells [34]. As illustrated by our study, integrating knowledge across
319 scales, i.e. going from the microbiological and immunological variables to individual
320 health is data-intensive and requires multidisciplinary knowledge from molecular biology
321 to public health. Nonetheless, it is a necessary and worthwhile challenge to bridge the
322 gap between molecular mechanisms and health outcomes.

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326 **Disclosure of Interests**

327 The authors declare the following financial interests/personal relationships which may be
328 considered as potential competing interests: TW serves on advisory boards for MesoScale
329 discovery (Merck) Sharp & Dohme.

330 **Contributions**

331 NT, CLM, NBo, and SA designed the study. NT, CLM, BE, BR, TK, IGB, and SA
332 designed the experiments. CB, VB, SGro, MR, and NBe performed experiments. CH,
333 JRa, and TW contributed reagents, materials, and analysis tools. CB, VB, SGra, SGro,
334 MR, MB, CG, VT, VF, CLM, JRe, IGB, MSe, NBo, and SA contributed to study design,
335 patient recruitment, and clinical data acquisition. NT, IBU, BE, BR, CS, TK, CLM,

336 TK and SA performed data analyses. NT, IBU, IGB, CLM, TK and SA wrote the initial
337 version of the manuscript. All authors approved the final version of the manuscript.

338 **Ethics**

339 The PAPCLEAR trial was promoted by the Centre Hospitalier Universitaire (CHU)
340 de Montpellier and has been approved by the Comité de Protection des Personnes
341 (CPP) Sud Méditerranée I on 11 May 2016 (CPP number 16 42, reference number ID
342 RCB 2016-A00712-49); by the Comité Consultatif sur le Traitement de l'Information en
343 matière de Recherche dans le domaine de la Santé on 12 July 2016 (reference number
344 16.504); and by the Commission Nationale Informatique et Libertés on 16 December
345 2016 (reference number MMS/ABD/AR1612278, decision number DR-2016-488). This
346 trial was authorised by the Agence Nationale de Sécurité du Médicament et des Produits
347 de Santé on 20 July 2016 (reference 20160072000007). The ClinicalTrials.gov identifier
348 is NCT02946346. All participants provided written informed consent.

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357 **Data availability**

358 Table 1, 2, 3, 4, Figures 1, 2, 3, Supplementary Tables S1, S2, S3, S4, S5, S6, S7, S8
359 and supplementary figures S1, S2, S3 have associated raw data. The data that support
360 the findings of this study are available from the corresponding author upon request, and
361 data are available in the Zenodo public repository (10.5281/zenodo.6913875).

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Table 1: **Key characteristics of participants included in the study.** *n* indicates the number of individuals, p-value refers to the outcome of a Kruskal-Wallis Rank Sum Test (`kruskal.test` function in R), IQR is the interquartile range, SD the standard deviation.

| | Tampon | Menstrual cup | p |
|---|-------------------------|------------------------|-------|
| n | 107 | 31 | |
| PAPCLEAR follow-up duration (days, median [IQR]) | 220.00 [116.50, 374.50] | 270.00 [74.00, 437.50] | 0.740 |
| Lifetime number of partners (median [IQR]) | 8.00 [4.00, 15.00] | 11.00 [4.00, 16.50] | 0.396 |
| Age (years, median [IQR]) | 21.00 [20.00, 23.00] | 23.00 [22.00, 24.00] | 0.021 |
| Age at menarchy (years, median [IQR]) | 13.00 [12.00, 14.00] | 13.00 [12.00, 14.00] | 0.695 |
| Age at sexual debut (years, median [IQR]) | 17 [15.00, 18.00] | 16 [15.00, 17.00] | 0.338 |
| Body Mass Index (BMI, median [IQR]) | 21.15 [19.78, 23.40] | 21.62 [20.56, 23.22] | 0.483 |
| Smoking (%) | | | 0.129 |
| No | 64 (59.8) | 24 (77.4) | |
| Occasionally | 15 (14.0) | 4 (12.9) | |
| Often | 28 (26.2) | 3 (9.7) | |
| Antibiotics (last two weeks) = Yes (%) | 7 (6.5) | 3 (9.7) | 0.842 |
| Lubricant (last two weeks) (%) | 14 (13.1) | 5 (16.1) | 0.891 |
| Intercourse with a regular partner (last two weeks) (%) | 63 (58.9) | 18 (58.1) | 1.000 |
| Intercourse with an occasional partner (last two weeks) (%) | 15 (14.0) | 4 (12.9) | 1.000 |
| Stress level (past week) (%) | | | 0.168 |
| 0 (Min) | 20 (18.7) | 5 (16.1) | |
| 1 | 45 (42.1) | 13 (41.9) | |
| 2 | 30 (28.0) | 13 (41.9) | |
| 3 (Max) | 12 (11.2) | 0 (0.0) | |
| Menses (last two weeks) (%) (Yes) | 58 (54.2) | 20 (64.5) | 0.416 |
| HPV vaccinated (%) | 55 (51.4) | 19 (61.3) | 0.443 |
| HPV positive (focal) (%) | 49 (45.8) | 10 (32.3) | 0.256 |
| HPV positive (multiple HPVs) (%) | 33 (30.8) | 9 (29.0) | 1.000 |

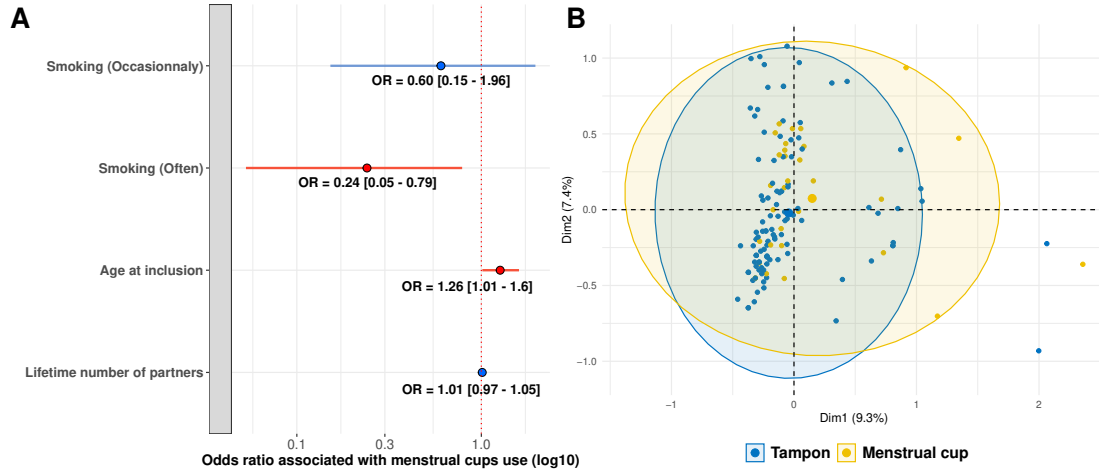


Figure 1: **Profile of menstrual cup users** A) Odds ratio of covariates associated with the use of menstrual cups. Results show the odds ratio (OR) of the factors selected in the best generalised linear model against menstrual cup as the response variable and using an Akaike Information Criterion corrected for small sample size (AICc). All covariates included in the model selection process are listed in Table S3. 95% confidence intervals are shown in brackets. B) Multiple Correspondence Analysis of socio-economic variables clustering participant using tampons or menstrual cups. Ellipses indicate 95% confidence intervals. Blue indicates tampon users, whereas yellow indicates women using menstrual cup.

Table 2: **Values of the main health metrics included in the study.** We show the number of occurrences and associated percentages (in parentheses) for each metric in the two study populations of size n . ‘p-value’ indicates the outcome of a t-test.

| | Tampon | Menstrual cup | p-value |
|--|-----------|---------------|---------|
| n | 107 | 31 | |
| Fungal genital infection (last 3 months) (%) | 7 (6.5) | 7 (22.6) | 0.023 |
| Genital infection (last 3 months) (%) | 5 (4.7) | 1 (3.2) | 1.000 |
| Chlamydia infection (last 3 months)(%) | 4 (3.7) | 1 (3.2) | 1.000 |
| Vaginosis (last 3 months)(%) | 1 (0.9) | 2 (6.5) | 0.248 |
| Urinary tract infection (last 3 months)(%) | 12 (11.2) | 6 (19.4) | 0.378 |
| HPV infection (focal) (%) | 49 (45.8) | 10 (32.3) | 0.256 |

Table 3: Factors associated with the reporting of a fungal genital infection. Results show the odds ratios (OR) of the factors selected in the best generalised linear model with fungal genital infection as the response variable and using an Akaike Information Criterion corrected for small sample size (AICc). All covariates included in the model selection process are listed in Table 4. SE stands for standard ‘error’. CI 2.5% and CI 97.5% indicate the lower and upper bounds of the 95% confidence interval.

| Response = Fungal genital infection | OR | OR SE | CI 2.5% | CI 97.5% | p.value |
|-------------------------------------|-------|-------|---------|----------|---------|
| Menstrual cup | 4.731 | 2.857 | 1.440 | 16.058 | 0.010 * |
| Lifetime number of partners | 0.998 | 0.027 | 0.938 | 1.047 | 0.952 |

* : $p < 0.05$

Table 4: Frequency of occurrence of the covariates among the 80 best GLM models with fungal genital infection as a response variable (Model 2). The models were selected using AICc, assuming an AICc + 5 cutoff. The columns indicate the number and percentage of models that contain each covariate, as well as the percentage of models where the associated p-value is marginally significant (lower than 10%) or significant (lower than 5%).

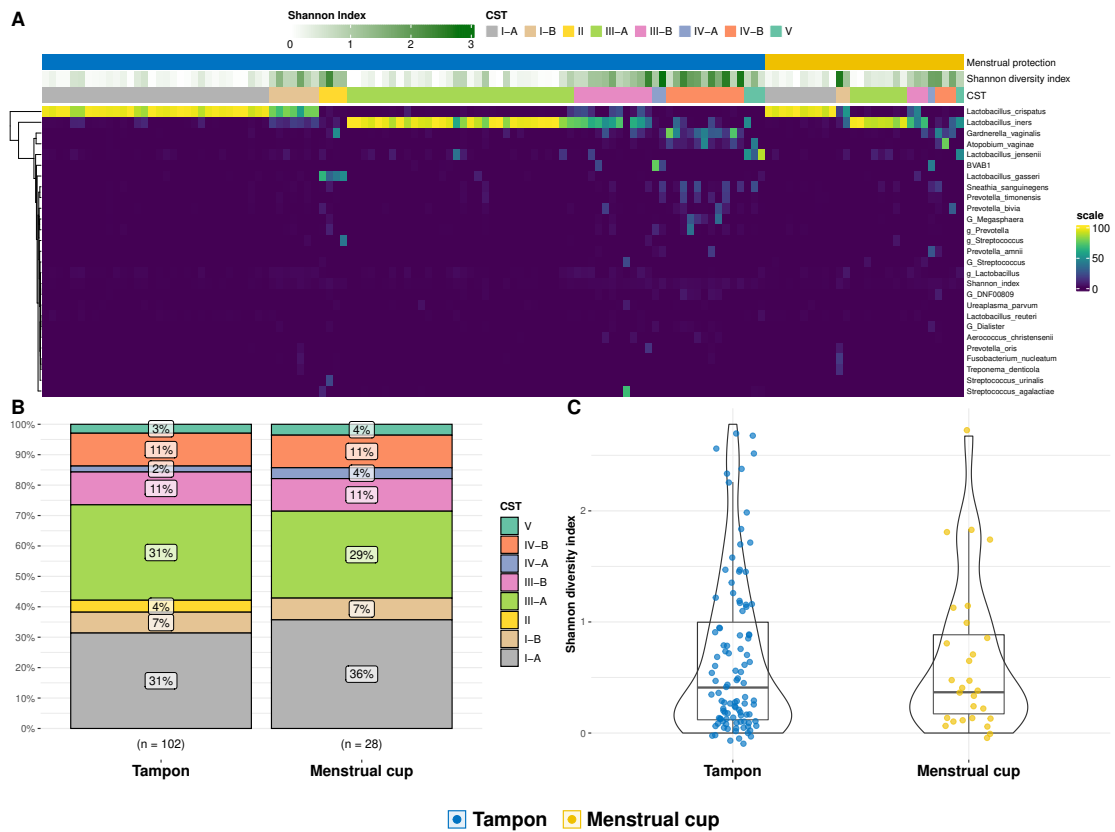


Figure 2: Metabarcoding differences between women using mainly menstrual cups or tampons. A) Abundance and diversity of the top 1% most abundant bacterial species found in participants, B) Community State Types (CST) distribution, C) Shannon diversity index. In A and B, colours show the type of menstrual product used (tampons in blue and cups in yellow).

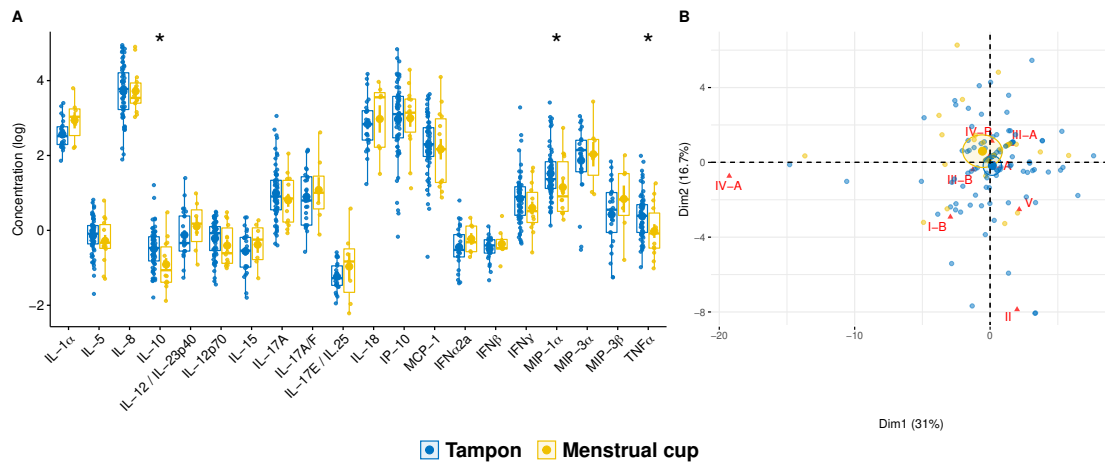


Figure 3: **Immunological differences between women using mainly menstrual cups or tampons.** A) Cytokines local concentrations (on a log scale), and B) Outcome of a multi-parametric clustering analysis using factor analysis of mixed data. In A, the presence of a star indicates significance ($p < 0.05$). Ellipses indicate 95% confidence intervals. In A and B, colours show the type of menstrual product used (tampons in blue and cups in yellow). In B, the vaginal Community State Types (CSTs) are shown in red.